

Confidence intervals

Announcements

- + HW 4 released, due next Friday ← *Note for Ciaran: modify this slightly*
- + Exam 1 released next Friday
 - + Take home, 1 week to complete
 - + Open note (anything from this course)
 - + Closed internet
 - + Closed other people
- + Reminder: department seminar on Monday, 12pm - 1pm (Dr. Mine Cetinkaya-Rundel)
 - + Can sign up to meet with the speaker 11 - 11:30

Wald vs. likelihood ratio tests

- Both can test one or more parameters
- Convention: use Wald for single parameters, LRT

typically $\nearrow$ for multiple

$$Z = \frac{\hat{\beta}_i - \beta_i^0}{\sqrt{\text{var}(\hat{\beta}_i)}} \sim N(0,1)$$

instead of $W = Z^2 \sim \chi^2_1$

- Wald tests do not require fitting a reduced model
- Wald tests are good for linear combinations of β s
- LRT does not require inverting (potentially large) matrices
- In logistic regression, LRT performs better when $\hat{p}_i \approx 0$ or 1 for some observations

95% CI : $1.192 \pm 1.96 (0.243) = (0.716, 1.668)$

Confidence intervals

↑
97.5 quantile of $N(0,1)$

$$\log\left(\frac{p_i}{1-p_i}\right) = \beta_0 + \beta_1 \text{Sex}_i + \beta_2 \text{Age}_i + \beta_3 \text{SecondClass}_i +$$

$\beta_4 \text{FirstClass}_i + \beta_5 \text{Sex}_i \cdot \text{Age}_i$
vcov(model)
...
Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
## (Intercept)	0.408232	0.330916	1.234	0.217337	
## Sexmale	-1.163444	0.437622	-2.659	0.007848	**
## Age	-0.007186	0.011684	-0.615	0.538522	
## Pclass2	1.191858	0.243233	4.900	9.58e-07	***
## Pclass1	2.697561	0.295822	9.119	< 2e-16	***
## Sexmale:Age	-0.049851	0.014782	-3.373	0.000745	***

...

diagonal entries of $\hat{\Sigma}^{-1}(\hat{\beta})$

How do I create a 95% confidence interval for β_3 ?

This statement is not true:

$$P(\beta_3 \in (0.716, 1.668)) = 0.95$$

This statement is true:

$$P(\beta_3 \in (\hat{\beta}_3 - 1.96 SE(\hat{\beta}_3), \hat{\beta}_3 + 1.96 SE(\hat{\beta}_3))) \\ = 0.95$$

Wald confidence intervals

Goal: want a "reasonable range" for unknown β_i

Method: $\hat{\beta} \approx N(\beta, \mathcal{I}^{-1}(\beta)) \Rightarrow \hat{\beta}_i \approx N(\beta_i, \underbrace{\mathcal{I}^{-1}(\beta)_{ii}}_{i^{\text{th}} \text{ diagonal entry } \mathcal{I}^{-1}(\beta)})$

Confidence interval: $\hat{\beta}_i \pm \underbrace{Z_{1-\frac{\alpha}{2}}}_{1-\frac{\alpha}{2} \text{ quantile of } N(0,1)} \underbrace{\sqrt{\mathcal{I}^{-1}(\hat{\beta})_{ii}}}_{SE(\hat{\beta}_i)}$

Ex: 95% CI $\Rightarrow \alpha = 0.05 \Rightarrow 1 - \frac{\alpha}{2} = 0.975$
 $Z_{1-\frac{\alpha}{2}} = 1.96 \Rightarrow \hat{\beta}_i \pm 1.96 \sqrt{\mathcal{I}^{-1}(\hat{\beta})_{ii}}$

CI's vs. Hypothesis tests

Let $\theta \in \mathbb{R}$ be a parameter of interest, and suppose

$\hat{\theta}$ is an estimate, where $\hat{\theta} \hat{\sim} N(\theta, \text{var}(\hat{\theta}))$

Consider testing $H_0: \theta = \theta_0$ vs. $H_A: \theta \neq \theta_0$

wald test statistic :
$$W = \frac{(\hat{\theta} - \theta_0)^2}{\text{var}(\hat{\theta})} \sim \chi^2_1$$

$$Z = \frac{\hat{\theta} - \theta_0}{\sqrt{\text{var}(\hat{\theta})}} \sim N(0, 1)$$

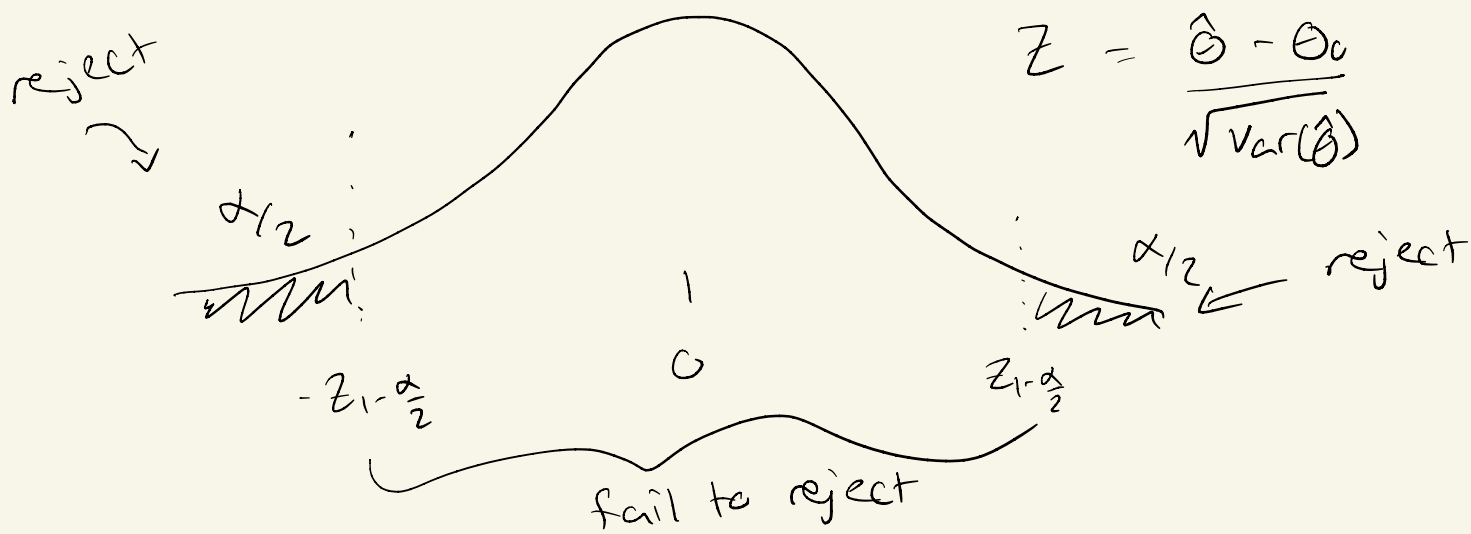
1) want a "reasonable range" of values for θ

2) Rejecting $H_0 \Rightarrow \theta = \theta_0$ is "unreasonable"

3) so, $CI = \{ \theta_0 : \text{fail to reject } H_0: \theta = \theta_0 \}$

Suppose we want type I error $= \alpha$, what is the rejection region?

$$Z < -Z_{1-\frac{\alpha}{2}} \quad \text{or} \quad Z > Z_{1-\frac{\alpha}{2}}$$



fail to reject when

$$-z_{1-\frac{\alpha}{2}} < \frac{\hat{\theta} - \theta_0}{\sqrt{\text{Var}(\hat{\theta})}} < z_{1-\frac{\alpha}{2}}$$

$$\Rightarrow \hat{\theta} - z_{1-\frac{\alpha}{2}} \sqrt{\text{Var}(\hat{\theta})} < \theta_0 < \hat{\theta} + z_{1-\frac{\alpha}{2}} \sqrt{\text{Var}(\hat{\theta})}$$

$$\Rightarrow CI = \left[\hat{\theta} - z_{1-\frac{\alpha}{2}} \sqrt{\text{Var}(\hat{\theta})}, \hat{\theta} + z_{1-\frac{\alpha}{2}} \sqrt{\text{Var}(\hat{\theta})} \right]$$

Technique: Inverting the test

Confidence intervals for linear combinations

Class activity

https://sta712-f22.github.io/class_activities/ca_lecture_15.html