# Goodness of fit and overdispersion

· No class on Friday (cut of town)

- I will post an activity or a pre-recorded lecture

· Hw 5 released

Recap: EDMs and deviance

$$f(y; \mu, \emptyset) = b(y, \emptyset) \exp \frac{2}{2} - \frac{\partial (y, \mu)}{2} = \frac{\partial (y, \mu)}{\partial y} =$$

Scaled residual deriance: 
$$D^*(y, \hat{n}) = \frac{D(y, \hat{n})}{\emptyset}$$

$$L(\beta) = \frac{1}{11} b(\lambda_i, \emptyset) \cdot \exp \left\{ -\frac{\partial(\lambda_i, \lambda_i)}{2\emptyset} \right\}$$

$$2 l(\beta) = 2 \frac{1}{2} \log b(\lambda_i, \emptyset) - \frac{1}{2} \frac{\partial(\lambda_i, \lambda_i)}{\emptyset}$$

$$3(\lambda_i) = \beta_i \lambda_i$$

$$g(\mu_i) = \beta^i \chi_i$$

$$\hat{\mu}_i = g^{-1}(\beta^i \chi_i) \qquad -0^*(y, \hat{\mu})$$

$$2 l(saturated) = 2 \hat{\xi} log b(\chi_i, \emptyset) - \hat{\xi} \frac{\partial(\chi_i, \chi_i)}{\partial \xi}$$

$$2 l(saturated) = 2 \frac{2}{i=1} log b(Vi, \emptyset) - \frac{2}{i=1} \frac{\partial (Vi, Vi)}{\partial i}$$

$$2 (l(saturated) - l(\hat{\beta})) = D^*(y, \hat{\mu})$$

Examples: tly, w) = ylogu - w > Poisson: convention; t(y,y) = ylogy - y 0 log 0 = 0  $\Rightarrow \partial(y,n) = 2\left(y\log\left(\frac{y}{n}\right) - (y-n)\right)$  $t(y,n) = y \log(\frac{n}{1-n}) + \log(1-n)$ 2) Bernaulli: = y log M + (1-y) log (1-w) try,y)= ylogy + (1-y) log(1-y) => d(y, M) = 2(y log ( ) + (1-y) by ( )  $t(y, w) = -\frac{1}{2}(y-w)^{2} = -\frac{1}{2}(y-y)^{2} = 0$   $= 7 d(y, w) = (y-w)^{2}$ 3) Norma): t(y,u) = yu - 1/2

V(m) = 2M Var(V) = Q. V(m) Norma); 1=(m)~ Saddlepoint approximation Paisson: Residual deviane: Dly, M) = Éd (11, Mi) ィピンこか want  $D(y, h) \propto \chi^2_{n-(u+1)}$  need distribution of Ollingi) Dispersion model form: b(y, Ø) exp\( \frac{-\partial(y, w)}{2\pi}\) Sadolepoint approximation; b(y, Ø) ~ 1/2110/(y) =>  $f(y; M, \emptyset) \sim \sqrt{2\pi g \nu(y)} \exp \left\{-\frac{\partial(y, M)}{2g}\right\}$ If subdispoint approximation holds,  $\frac{\partial(Y_i, u_i)}{\partial X_i} \approx X_i^2$  (S.4.3 in book)  $= 7 D^*(y, u) = \frac{1}{2} \frac{2}{2} d(Y_i, u_i) \approx X_i^2$  [Intrition: lose 1 of for each estimated parameter parameter (cochron's theorem) (cochron's theorem) 3/14 Scallepoint approximation · Normal: "approximation is exact · Paisson:

indicates lack of fit

want minstiz 23 · Bernoulli : approximation doesn't hold

=> no goodness of fit test for binary response

For Poisson data, if min { 1;3 43

tends to be conserative, so a small produce still

the GOF test

#### **Data**

A concerned parent asks us to investigate crime rates on college campuses. We have access to data on 81 different colleges and universities in the US, including the following variables:

- type: college (C) or university (U)
- nv: the number of violent crimes for that institution in the given year
- enroll1000: the number of enrolled students, in thousands
- region: region of the US C = Central, MW = Midwest, NE = Northeast, SE = Southeast, SW = Southwest, and W = West)

## Goodness of fit

Scaled

Goodness of fit test: If the model is a good fit for the data, then the residual deviance follows a  $\chi^2$  distribution with the same degrees of freedom as the residual deviance

Residual deviance = 621.24, df = 75

So our model might not be a very good fit to the data.

Why might our model not be a good fit?

#### **Offsets**

We will account for school size by including an **offset** in the model:

$$\log(\lambda_i) = eta_0 + eta_1 M W_i + eta_2 N E_i + eta_3 S E_i + eta_4 S W_i + eta_5 W_i + \log(Enrollment_i)$$

offset term

where no Bi.

### **Motivation for offsets**

We can rewrite our regression model with the offset:

$$egin{aligned} \log(\lambda_i) &= eta_0 + eta_1 M W_i + eta_2 N E_i + eta_3 S E_i + eta_4 S W_i + eta_5 W_i \ &+ \log(Enrollment_i) \end{aligned}$$

# Fitting a model with an offset

The offset doesn't show up in the output (because we're not estimating a coefficient for it)

# Fitting a model with an offset

$$egin{split} \log(\widehat{\lambda}_i) &= -1.30 + 0.10 MW_i + 0.76 NE_i + \ 0.87 SE_i + 0.51 SW_i + 0.21 W_i \ &+ \log(Enrollment_i) \end{split}$$

How would I interpret the intercept -1.30?

### When to use offsets

Offsets are useful in Poisson regression when our counts come from groups of very different sizes (e.g., different numbers of students on a college campus). The offset lets us interpret model coefficients in terms of rates instead of raw counts.

With your neighbor, brainstorm some other data scenarios where our response is a count variable, and an offset would be useful. What would our offset be?

#### Goodness of fit

```
m2 <- glm(nv ~ region, offset = log(enroll1000),
          data = crimes, family = poisson)
summary(m2)
## (Dispersion parameter for poisson family taken to be 1)
##
      Null deviance: 491.00 on 80 degrees of freedom
##
## Residual deviance: 433.14 on 75 degrees of freedom
pchisq(433.14, df=75, lower.tail=F)
## [1] 8.33082e-52
```

# Overdispersion

Overdispersion occurs when the response Y has higher variance than we would expect from the specified EDM

Why is it a problem if Y has more variance than we account for in our model?

# Estimating $\phi$

# Using $\widehat{\phi}$

```
pearson_resids <- residuals(m2, type="pearson")</pre>
sum(pearson resids^2)/df.residual(m2)
## [1] 7.58542
             Estimate Std. Error z value Pr(>|z|)
##
                        0.12403 -10.517 < 2e-16 ***
##
  (Intercept) -1.30445
  regionMW 0.09754
                        0.17752 0.549 0.58270
## regionNE 0.76268 0.15292 4.987 6.12e-07 ***
## regionSE 0.87237
                        0.15313 5.697 1.22e-08 ***
## regionSW 0.50708
                        0.18507 2.740 0.00615 **
## regionW
                        0.18605 1.125 0.26053
          0.20934
```